## Tentamentsskrivning i Statistisk slutledning MVE155/MSG200, 7.5 hp .

Tid: 17 mars 2020, kl 14.00-18.00
Examinator och jour: Serik Sagitov, tel. 031-772-5351
CTH: för " 3 " fordras 12 poäng, för " 4 " - 18 poäng, för " 5 " - 24 poäng. GU : för "G" fordras 12 poäng, för "VG" - 20 poäng.
Inclusive eventuella bonuspoäng.
This is a home-written exam.
The examination must be conducted individually, that is, cooperation is not allowed.
You have 4 hours to complete the exam. Solutions are written on paper, or digitally on a digital writing pad if you have access to it. Never write more than one task on each sheet. After 4 hours, you have 30 minutes to scan / photograph your solutions and organize and submit your solutions according to one of the following ways in order of priority:

1. A single pdf file where the pages are arranged in the order of the questions.
2. Image files (jpg or png) or pdf files where each file contains solution to just one question and named according to "Question 1", or "Question 1 page 1", Question 1 page 2 "etc if there are multiple pages to a question.
3. (5 points) Miscellaneous questions.
(a) The Oxford English Dictionary is widely regarded as the accepted authority on the English language. It is an unsurpassed guide to the meaning, history, and pronunciation of 600,000 words, past and present, from across the English-speaking world. Propose a statistical inference algorithm for estimating your English vocabulary size using the online version of the Oxford English Dictionary.
(b) Explain the difference between the Kruskal-Wallis and the Fridman tests referring to the dataset

| Placebo | Treatment 1 | Treatment 2 |
| :--- | :--- | :--- |
| 174 | 263 | 105 |
| 224 | 213 | 103 |
| 260 | 231 | 145 |
| 225 | 291 | 108 |

(c) Given that a beta posterior distribution is skewed to the right, which of the two estimates for the population proportion would be larger: the MAP estimate or the PME estimate? Explain by drawing a skewed posterior curve.
2. ( 5 points) A population consists of three subpopulations whose relative sizes are $w_{1}=0.5$, $w_{2}=0.25, w_{3}=0.25$. The variable of interest $X$, characterising a generic element of the population, is normally distributed over each of the subpopulations. The normal distribution of subpopulation $i$ has mean $\mu_{i}$ (possibly different for different $i=1,2,3$ ) and standard deviation $\sigma$ which is the same across subpopulations.
(a) In the current setting, what is the optimal allocation of 100 sample observations among three strata for estimation of the population mean $\mu$ ? Explain.
(b) Three independent samples each of size 10 were drawn from the three strata. The data produced three sample means $\bar{x}_{1}=-0.3, \bar{x}_{2}=0.8, \bar{x}_{3}=-0.5$ and the pooled sample standard deviation $s_{p}=1.1$. Do the data reveal a significant difference among three strata means $\mu_{1}, \mu_{2}, \mu_{3}$ ?
(c) Suppose we know that $\mu_{1}=0, \mu_{2}=1, \mu_{3}=2$, and $\sigma=1$. What is the variance $\operatorname{Var}(X)$ for the whole population in this case?
(d) Given $\mu_{1}=0, \mu_{2}=1, \mu_{3}=2$, and $\sigma=1$, draw a sketch depicting three subpopulation distributions. On top of these three curves draw the overall population distribution.
3. ( 5 marks) A computer program has simulated two independent values $\left(x_{1}, x_{2}\right)$ from a normal distribution $\mathrm{N}(0, \sigma)$ with zero mean and unknown to you standard deviation $\sigma$.
(a) Show that

$$
\hat{\sigma}^{2}=\frac{x_{1}^{2}+x_{2}^{2}}{2}
$$

is an unbiased estimate of the variance $\sigma^{2}$. Would this estimate be biased if the random variables $\left(X_{1}, X_{2}\right)$ are positively correlated?
(b) The scaled estimate

$$
\frac{2 \hat{\sigma}^{2}}{\sigma^{2}}
$$

has a particular sampling distribution, what is it? Using the table for this distribution, construct a formula for an exact $95 \%$ confidence interval for $\sigma^{2}$.
(c) Show that

$$
\tilde{\sigma}^{2}=\frac{\left(x_{1}-x_{2}\right)^{2}}{2}
$$

is also is an unbiased estimate of the variance $\sigma^{2}$. What its relation to the sample variance $s^{2}$ ? Would this estimate be biased if the random variables $\left(X_{1}, X_{2}\right)$ are positively correlated?
(d) Which of these two unbiased point estimates of $\sigma^{2}$ would you prefer? Explain why.
4. (5 points) The following (ordered) 16 numbers are computer generated from $\mathrm{N}(\mu, \sigma)$

| 0.11 | 1.60 | 1.61 | 1.72 |
| :--- | :--- | :--- | :--- |
| 2.28 | 3.12 | 3.15 | 3.53 |
| 3.70 | 4.15 | 4.25 | 4.74 |
| 5.33 | 5.49 | 6.39 | 6.59 |

(a) Compute the inter-quartile range for this sample. In what sense this measure of dispersion is robust against outliers?
(b) Estimate $\sigma$ using the the inter-quartile range of the standard normal distribution.
(c) Sketch the normal probability plot using only lower quartiles, medians, and upper quartiles. Explain how you did it step by step.
5. (5 points) In a political poll survey two independently chosen at random groups of voters were asked whether they would vote for the political party $\mathcal{P}$. Each person answered either yes or no (which even includes don't know option). Group 2, which consisted of 2000 voters, was contacted in August 2019, while group 1, which also consisted of 2000 voters, was asked first in August 2018 and then again in August 2019. The purpose of the survey was to compare two population proportions:

- $p_{1}$ proportion of people supporting party $\mathcal{P}$ in August 2018,
- $p_{2}$ proportion of people supporting party $\mathcal{P}$ in August 2019.

The percentages of yes answers obtained by the survey were as follows

- group 1 in August 2018: 10\%,
- group 1 in August 2019: 12\%,
- group 2 in August 2019: 13\%.
(a) What is you best point estimate of the population proportion $p_{2}$. Compute its standard error.
(b) Find a $95 \%$ confidence interval for the difference $p_{2}-p_{1}$. Justify the choice of the formula you apply.
(c) Using the additional information that 200 out 2000 people in the group 1 have changed their answers between 2018 and 2019 (either from yes to no, or from no to yes), and disregarding the response of group 2, test $H_{0}: p_{1}=p_{2}$ against $H_{1}: p_{1} \neq p_{2}$.

6 (5 marks) The data below show rounded-to-integer values of $x=$ frequency ( MHz ) and $y=$ output power (W) for a certain laser configuration.

$$
\begin{array}{r|rrrrrrrr}
x & 60 & 63 & 77 & 100 & 125 & 157 & 186 & 222 \\
\hline y & 16 & 17 & 19 & 21 & 22 & 20 & 15 & 5
\end{array}
$$

The Matlab 'regress' command yields the following information for a quadratic regression model:

```
b}
    -1.5127
    0.3919
    -0.0016
bint =
    -2.9440-0.0814
    0.3678 0.4160
    -0.0017 -0.0015
r =
    -0.1283 0.2980 0.0089-0.3634 0.0158 0.1968 0.0593-0.0870
```

where b stands for the estimated parameters $\beta_{0}, \beta_{1}, \beta_{2}$, bint gives three $95 \%$ confidents intervals, and $r$ gives 8 residuals. The sum of squares of the residuals is 0.2874 . The sample standard deviation of $y$ is 5.3835 .
(a) Does the quadratic model appear to be suitable for explaining observed variation in output power by relating it to frequency? Answer by applying a relevant parametric statistical test. What are your assumptions about the underlying statistical model? How do you verify the key assumption?
(b) Find the adjusted coefficient of determination. What does is say?
(c) The sum of the 8 residuals is 0 . Prove in the simple linear regression setting that the sum of the residuals equals zero.
(d) Draw by hand the scatter plot for the data and then on top of the scatter plot draw the line predicted by the quadratic model.

Normal distribution table

| $z$ | 0.00 | 0.01 | 0.02 | 0.03 | 0.04 | 0.05 | 0.06 | 0.07 | 0.08 | 0.09 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.0 | 0.5000 | 0.5040 | 0.5080 | 0.5120 | 0.5160 | 0.5199 | 0.5239 | 0.5279 | 0.5319 | 0.5359 |
| 0.1 | 0.5398 | 0.5438 | 0.5478 | 0.5517 | 0.5557 | 0.5596 | 0.5636 | 0.5675 | 0.5714 | 0.5753 |
| 0.2 | 0.5793 | 0.5832 | 0.5871 | 0.5910 | 0.5948 | 0.5987 | 0.6026 | 0.6064 | 0.6103 | 0.6141 |
| 0.3 | 0.6179 | 0.6217 | 0.6255 | 0.6293 | 0.6331 | 0.6368 | 0.6406 | 0.6443 | 0.6480 | 0.6517 |
| 0.4 | 0.6554 | 0.6591 | 0.6628 | 0.6664 | 0.6700 | 0.6736 | 0.6772 | 0.6808 | 0.6844 | 0.6879 |
| 0.5 | 0.6915 | 0.6950 | 0.6985 | 0.7019 | 0.7054 | 0.7088 | 0.7123 | 0.7157 | 0.7190 | 0.7224 |
| 0.6 | 0.7257 | 0.7291 | 0.7324 | 0.7357 | 0.7389 | 0.7422 | 0.7454 | 0.7486 | 0.7517 | 0.7549 |
| 0.7 | 0.7580 | 0.7611 | 0.7642 | 0.7673 | 0.7704 | 0.7734 | 0.7764 | 0.7794 | 0.7823 | 0.7852 |
| 0.8 | 0.7881 | 0.7910 | 0.7939 | 0.7967 | 0.7995 | 0.8023 | 0.8051 | 0.8078 | 0.8106 | 0.8133 |
| 0.9 | 0.8159 | 0.8186 | 0.8212 | 0.8238 | 0.8264 | 0.8289 | 0.8315 | 0.8340 | 0.8365 | 0.8389 |
| 1.0 | 0.8413 | 0.8438 | 0.8461 | 0.8485 | 0.8508 | 0.8531 | 0.8554 | 0.8577 | 0.8599 | 0.8621 |
| 1.1 | 0.8643 | 0.8665 | 0.8686 | 0.8708 | 0.8729 | 0.8749 | 0.8770 | 0.8790 | 0.8810 | 0.8830 |
| 1.2 | 0.8849 | 0.8869 | 0.8888 | 0.8907 | 0.8925 | 0.8944 | 0.8962 | 0.8980 | 0.8997 | 0.9015 |
| 1.3 | 0.9032 | 0.9049 | 0.9066 | 0.9082 | 0.9099 | 0.9115 | 0.9131 | 0.9147 | 0.9162 | 0.9177 |
| 1.4 | 0.9192 | 0.9207 | 0.9222 | 0.9236 | 0.9251 | 0.9265 | 0.9279 | 0.9292 | 0.9306 | 0.9319 |
| 1.5 | 0.9332 | 0.9345 | 0.9357 | 0.9370 | 0.9382 | 0.9394 | 0.9406 | 0.9418 | 0.9429 | 0.9441 |
| 1.6 | 0.9452 | 0.9463 | 0.9474 | 0.9484 | 0.9495 | 0.9505 | 0.9515 | 0.9525 | 0.9535 | 0.9545 |
| 1.7 | 0.9554 | 0.9564 | 0.9573 | 0.9582 | 0.9591 | 0.9599 | 0.9608 | 0.9616 | 0.9625 | 0.9633 |
| 1.8 | 0.9641 | 0.9649 | 0.9656 | 0.9664 | 0.9671 | 0.9678 | 0.9686 | 0.9693 | 0.9699 | 0.9706 |
| 1.9 | 0.9713 | 0.9719 | 0.9726 | 0.9732 | 0.9738 | 0.9744 | 0.9750 | 0.9756 | 0.9761 | 0.9767 |
| 2.0 | 0.9772 | 0.9778 | 0.9783 | 0.9788 | 0.9793 | 0.9798 | 0.9803 | 0.9808 | 0.9812 | 0.9817 |
| 2.1 | 0.9821 | 0.9826 | 0.9830 | 0.9834 | 0.9838 | 0.9842 | 0.9846 | 0.9850 | 0.9854 | 0.9857 |
| 2.2 | 0.9861 | 0.9864 | 0.9868 | 0.9871 | 0.9875 | 0.9878 | 0.9881 | 0.9884 | 0.9887 | 0.9890 |
| 2.3 | 0.9893 | 0.9896 | 0.9898 | 0.9901 | 0.9904 | 0.9906 | 0.9909 | 0.9911 | 0.9913 | 0.9916 |
| 2.4 | 0.9918 | 0.9920 | 0.9922 | 0.9925 | 0.9927 | 0.9929 | 0.9931 | 0.9932 | 0.9934 | 0.9936 |
| 2.5 | 0.9938 | 0.9940 | 0.9941 | 0.9943 | 0.9945 | 0.9946 | 0.9948 | 0.9949 | 0.9951 | 0.9952 |
| 2.6 | 0.9953 | 0.9955 | 0.9956 | 0.9957 | 0.9959 | 0.9960 | 0.9961 | 0.9962 | 0.9963 | 0.9964 |
| 2.7 | 0.9965 | 0.9966 | 0.9967 | 0.9968 | 0.9969 | 0.9970 | 0.9971 | 0.9972 | 0.9973 | 0.9974 |
| 2.8 | 0.9974 | 0.9975 | 0.9976 | 0.9977 | 0.9977 | 0.9978 | 0.9979 | 0.9979 | 0.9980 | 0.9981 |
| 2.9 | 0.9981 | 0.9982 | 0.9982 | 0.9983 | 0.9984 | 0.9984 | 0.9985 | 0.9985 | 0.9986 | 0.9986 |
| 3.0 | 0.9987 | 0.9987 | 0.9987 | 0.9988 | 0.9988 | 0.9989 | 0.9989 | 0.9989 | 0.9990 | 0.9990 |
| 3.1 | 0.9990 | 0.9991 | 0.9991 | 0.9991 | 0.9992 | 0.9992 | 0.9992 | 0.9992 | 0.9993 | 0.9993 |
| 3.2 | 0.9993 | 0.9993 | 0.9994 | 0.9994 | 0.9994 | 0.9994 | 0.9994 | 0.9995 | 0.9995 | 0.9995 |
| 3.3 | 0.9995 | 0.9995 | 0.9995 | 0.9996 | 0.9996 | 0.9996 | 0.9996 | 0.9996 | 0.9996 | 0.9997 |
| 3.4 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9998 |

Chi-square distribution table


Area to the Right of the Critical Value of $\chi^{2}$

| df | 0.995 | 0.990 | 0.975 | 0.950 | 0.900 | 0.100 | 0.050 | 0.025 | 0.010 | 0.005 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 0.000 | 0.000 | 0.001 | 0.004 | 0.016 | 2.706 | 3.841 | 5.024 | 6.635 | 7.879 |
| 2 | 0.010 | 0.020 | 0.051 | 0.103 | 0.211 | 4.605 | 5.991 | 7.378 | 9.210 | 10.597 |
| 3 | 0.072 | 0.115 | 0.216 | 0.352 | 0.584 | 6.251 | 7.815 | 9.348 | 11.345 | 12.838 |
| 4 | 0.207 | 0.297 | 0.484 | 0.711 | 1.064 | 7.779 | 9.488 | 11.143 | 13.277 | 14.860 |
| 5 | 0.412 | 0.554 | 0.831 | 1.145 | 1.610 | 9.236 | 11.070 | 12.833 | 15.086 | 16.750 |
| 6 | 0.676 | 0.872 | 1.237 | 1.635 | 2.204 | 10.645 | 12.592 | 14.449 | 16.812 | 18.548 |
| 7 | 0.989 | 1.239 | 1.690 | 2.167 | 2.833 | 12.017 | 14.067 | 16.013 | 18.475 | 20.278 |
| 8 | 1.344 | 1.646 | 2.180 | 2.733 | 3.490 | 13.362 | 15.507 | 17.535 | 20.090 | 21.955 |
| 9 | 1.735 | 2.088 | 2.700 | 3.325 | 4.168 | 14.684 | 16.919 | 19.023 | 21.666 | 23.589 |
| 10 | 2.156 | 2.558 | 3.247 | 3.940 | 4.865 | 15.987 | 18.307 | 20.483 | 23.209 | 25.188 |
| 11 | 2.603 | 3.053 | 3.816 | 4.575 | 5.578 | 17.275 | 19.675 | 21.920 | 24.725 | 26.757 |
| 12 | 3.074 | 3.571 | 4.404 | 5.226 | 6.304 | 18.549 | 21.026 | 23.337 | 26.217 | 28.300 |
| 13 | 3.565 | 4.107 | 5.009 | 5.892 | 7.042 | 19.812 | 22.362 | 24.736 | 27.688 | 29.819 |
| 14 | 4.075 | 4.660 | 5.629 | 6.571 | 7.790 | 21.064 | 23.685 | 26.119 | 29.141 | 31.319 |
| 15 | 4.601 | 5.229 | 6.262 | 7.261 | 8.547 | 22.307 | 24.996 | 27.488 | 30.578 | 32.801 |
| 16 | 5.142 | 5.812 | 6.908 | 7.962 | 9.312 | 23.542 | 26.296 | 28.845 | 32.000 | 34.267 |
| 17 | 5.697 | 6.408 | 7.564 | 8.672 | 10.085 | 24.769 | 27.587 | 30.191 | 33.409 | 35.718 |
| 18 | 6.265 | 7.015 | 8.231 | 9.390 | 10.865 | 25.989 | 28.869 | 31.526 | 34.805 | 37.156 |
| 19 | 6.844 | 7.633 | 8.907 | 10.117 | 11.651 | 27.204 | 30.144 | 32.852 | 36.191 | 38.582 |
| 20 | 7.434 | 8.260 | 9.591 | 10.851 | 12.443 | 28.412 | 31.410 | 34.170 | 37.566 | 39.997 |
| 21 | 8.034 | 8.897 | 10.283 | 11.591 | 13.240 | 29.615 | 32.671 | 35.479 | 38.932 | 41.401 |
| 22 | 8.643 | 9.542 | 10.982 | 12.338 | 14.041 | 30.813 | 33.924 | 36.781 | 40.289 | 42.796 |
| 23 | 9.260 | 10.196 | 11.689 | 13.091 | 14.848 | 32.007 | 35.172 | 38.076 | 41.638 | 44.181 |
| 24 | 9.886 | 10.856 | 12.401 | 13.848 | 15.659 | 33.196 | 36.415 | 39.364 | 42.980 | 45.559 |
| 25 | 10.520 | 11.524 | 13.120 | 14.611 | 16.473 | 34.382 | 37.652 | 40.646 | 44.314 | 46.928 |
| 26 | 11.160 | 12.198 | 13.844 | 15.379 | 17.292 | 35.563 | 38.885 | 41.923 | 45.642 | 48.290 |
| 27 | 11.808 | 12.879 | 14.573 | 16.151 | 18.114 | 36.741 | 40.113 | 43.195 | 46.963 | 49.645 |
| 28 | 12.461 | 13.565 | 15.308 | 16.928 | 18.939 | 37.916 | 41.337 | 44.461 | 48.278 | 50.993 |
| 29 | 13.121 | 14.256 | 16.047 | 17.708 | 19.768 | 39.087 | 42.557 | 45.722 | 49.588 | 52.336 |
| 30 | 13.787 | 14.953 | 16.791 | 18.493 | 20.599 | 40.256 | 43.773 | 46.979 | 50.892 | 53.672 |
| 40 | 20.707 | 22.164 | 24.433 | 26.509 | 29.051 | 51.805 | 55.758 | 59.342 | 63.691 | 66.766 |
| 50 | 27.991 | 29.707 | 32.357 | 34.764 | 37.689 | 63.167 | 67.505 | 71.420 | 76.154 | 79.490 |
| 60 | 35.534 | 37.485 | 40.482 | 43.188 | 46.459 | 74.397 | 79.082 | 83.298 | 88.379 | 91.952 |
| 70 | 43.275 | 45.442 | 48.758 | 51.739 | 55.329 | 85.527 | 90.531 | 95.023 | 100.425 | 104.215 |
| 80 | 51.172 | 53.540 | 57.153 | 60.391 | 64.278 | 96.578 | 101.879 | 106.629 | 112.329 | 116.321 |
| 90 | 59.196 | 61.754 | 65.647 | 69.126 | 73.291 | 107.565 | 113.145 | 118.136 | 124.116 | 128.299 |
| 100 | 67.328 | 70.065 | 74.222 | 77.929 | 82.358 | 118.498 | 124.342 | 129.561 | 135.807 | 140.169 |

Critical values of $t$-distribution

| df/ $\boldsymbol{\alpha}=$ | .40 | .25 | .10 | .05 | .025 | .01 | .005 | .001 | .0005 |
| ---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | ---: |
| 1 | 0.325 | 1.000 | 3.078 | 6.314 | 12.706 | 31.821 | 63.657 | 318.309 | 636.619 |
| 2 | 0.289 | 0.816 | 1.886 | 2.920 | 4.303 | 6.965 | 9.925 | 22.327 | 31.599 |
| 3 | 0.277 | 0.765 | 1.638 | 2.353 | 3.182 | 4.541 | 5.841 | 10.215 | 12.924 |
| 4 | 0.271 | 0.741 | 1.533 | 2.132 | 2.776 | 3.747 | 4.604 | 7.173 | 8.610 |
| 5 | 0.267 | 0.727 | 1.476 | 2.015 | 2.571 | 3.365 | 4.032 | 5.893 | 6.869 |
| 6 | 0.265 | 0.718 | 1.440 | 1.943 | 2.447 | 3.143 | 3.707 | 5.208 | 5.959 |
| 7 | 0.263 | 0.711 | 1.415 | 1.895 | 2.365 | 2.998 | 3.499 | 4.785 | 5.408 |
| 8 | 0.262 | 0.706 | 1.397 | 1.860 | 2.306 | 2.896 | 3.355 | 4.501 | 5.041 |
| 9 | 0.261 | 0.703 | 1.383 | 1.833 | 2.262 | 2.821 | 3.250 | 4.297 | 4.781 |
| 10 | 0.260 | 0.700 | 1.372 | 1.812 | 2.228 | 2.764 | 3.169 | 4.144 | 4.587 |
| 11 | 0.260 | 0.697 | 1.363 | 1.796 | 2.201 | 2.718 | 3.106 | 4.025 | 4.437 |
| 12 | 0.259 | 0.695 | 1.356 | 1.782 | 2.179 | 2.681 | 3.055 | 3.930 | 4.318 |
| 13 | 0.259 | 0.694 | 1.350 | 1.771 | 2.160 | 2.650 | 3.012 | 3.852 | 4.221 |
| 14 | 0.258 | 0.692 | 1.345 | 1.761 | 2.145 | 2.624 | 2.977 | 3.787 | 4.140 |
| 15 | 0.258 | 0.691 | 1.341 | 1.753 | 2.131 | 2.602 | 2.947 | 3.733 | 4.073 |
| 16 | 0.258 | 0.690 | 1.337 | 1.746 | 2.120 | 2.583 | 2.921 | 3.686 | 4.015 |
| 17 | 0.257 | 0.689 | 1.333 | 1.740 | 2.110 | 2.567 | 2.898 | 3.646 | 3.965 |
| 18 | 0.257 | 0.688 | 1.330 | 1.734 | 2.101 | 2.552 | 2.878 | 3.610 | 3.922 |
| 19 | 0.257 | 0.688 | 1.328 | 1.729 | 2.093 | 2.539 | 2.861 | 3.579 | 3.883 |
| 20 | 0.257 | 0.687 | 1.325 | 1.725 | 2.086 | 2.528 | 2.845 | 3.552 | 3.850 |
| 21 | 0.257 | 0.686 | 1.323 | 1.721 | 2.080 | 2.518 | 2.831 | 3.527 | 3.819 |
| 22 | 0.256 | 0.686 | 1.321 | 1.717 | 2.074 | 2.508 | 2.819 | 3.505 | 3.792 |
| 23 | 0.256 | 0.685 | 1.319 | 1.714 | 2.069 | 2.500 | 2.807 | 3.485 | 3.768 |
| 24 | 0.256 | 0.685 | 1.318 | 1.711 | 2.064 | 2.492 | 2.797 | 3.467 | 3.745 |
| 25 | 0.256 | 0.684 | 1.316 | 1.708 | 2.060 | 2.485 | 2.787 | 3.450 | 3.725 |
| 26 | 0.256 | 0.684 | 1.315 | 1.706 | 2.056 | 2.479 | 2.779 | 3.435 | 3.707 |
| 27 | 0.256 | 0.684 | 1.314 | 1.703 | 2.052 | 2.473 | 2.771 | 3.421 | 3.690 |
| 28 | 0.256 | 0.683 | 1.313 | 1.701 | 2.048 | 2.467 | 2.763 | 3.408 | 3.674 |
| 29 | 0.256 | 0.683 | 1.311 | 1.699 | 2.045 | 2.462 | 2.756 | 3.396 | 3.659 |
| 30 | 0.256 | 0.683 | 1.310 | 1.697 | 2.042 | 2.457 | 2.750 | 3.385 | 3.646 |
| 35 | 0.255 | 0.682 | 1.306 | 1.690 | 2.030 | 2.438 | 2.724 | 3.340 | 3.591 |
| 40 | 0.255 | 0.681 | 1.303 | 1.684 | 2.021 | 2.423 | 2.704 | 3.307 | 3.551 |
| 50 | 0.255 | 0.679 | 1.299 | 1.676 | 2.009 | 2.403 | 2.678 | 3.261 | 3.496 |
| 60 | 0.254 | 0.679 | 1.296 | 1.671 | 2.000 | 2.390 | 2.660 | 3.232 | 3.460 |
| 120 | 0.254 | 0.677 | 1.289 | 1.658 | 1.980 | 2.358 | 2.617 | 3.160 | 3.373 |
| inf. | 0.253 | 0.674 | 1.282 | 1.645 | 1.960 | 2.326 | 2.576 | 3.090 | 3.291 |
|  |  |  |  |  |  |  |  |  |  |

Critical values of the F-distribution (continued)

|  |  |  |  |  |  | Degrees of freedom in the numerator |  |  |  |  |
| :---: | :---: | ---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | 1 | 2 | 3 | 4 | 5 | 6 |

## NUMERICAL ANSWERS

1a. Choose at random $n=1000$ English words from the dictionary and compute the proportion $\hat{p}$ of the words you knew. A good point estimate of your English vocabulary size would be

$$
\hat{N}=\hat{p} \cdot 600000
$$

1b. The assumption for the Kruskal-Wallis is three independent samples and you rank the pooled sample

| Placebo | Treatment 1 | Treatment 2 |
| :--- | :--- | :--- |
| 5 | 11 | 2 |
| 7 | 6 | 1 |
| 10 | 9 | 4 |
| 8 | 12 | 3 |

and the main idea is to check if the three rank averages are significantly different.
For the Fridman test the three columns are not independent since they involve four subjects responding to the three treatments. Here the ranking is done separately by each subject

| Placebo | Treatment 1 | Treatment 2 |
| :--- | :--- | :--- |
| 2 | 3 | 1 |
| 3 | 2 | 1 |
| 3 | 2 | 1 |
| 2 | 3 | 1 |

and the Fridman idea is (again) to see if the rank means for treatments are significantly different.
1c. Typically, a distribution skewed to the right has a mean larger than its mode. In this case, PME > MAP.

2a. Given the three standard deviations are equal, the optimal allocation is the same as the proportional

$$
n_{1}=50, n_{2}=25, n_{3}=25
$$

2b. We apply the one-way anova test. Using the the samples means we obtain

$$
\begin{aligned}
& \bar{x}_{. .}=(10 \cdot(-0.3)+10 \cdot(0.8)+10 \cdot(-0.3)) / 30=0 \\
& S S_{A}=10 \cdot\left((0.3)^{2}+(0.8)^{2}+(0.3)^{2}\right)=9.8 \\
& M S_{A}=9.8 / 2=4.9
\end{aligned}
$$

Moreover, we know that $M S_{E}=s_{p}^{2}=1.21$. Thus the observed value of the F-test statistics is

$$
F=\frac{M S_{A}}{M S_{E}}=\frac{4.9}{1.21}=4.05 .
$$

Turning to the table for $F_{2,27}$ we find that the p-value of the test is between $2.5 \%$ and $5 \%$. We conclude that the difference between the three population means is significant at $5 \%$ significance level.

2c. Using

$$
\begin{aligned}
\mu & =w_{1} \mu_{1}+w_{2} \mu_{2}+w_{3} \mu_{3}=0.75, \\
\sigma^{2} & =\overline{\sigma^{2}}+\sum_{j=1}^{3} w_{j}\left(\mu_{j}-\mu\right)^{2},
\end{aligned}
$$

where

$$
\overline{\sigma^{2}}=w_{1} \sigma_{1}^{2}+w_{2} \sigma_{2}^{2}+w_{3} \sigma_{3}^{2}=1
$$

we find

$$
\sigma^{2}=1.69
$$

2d.


3a. Since $\mu=0$, we have

$$
\mathrm{E}\left(X_{i}^{2}\right)=\operatorname{Var}\left(X_{1}\right)=\sigma^{2}, \quad i=1,2
$$

and

$$
\mathrm{E}\left(\frac{X_{1}^{2}+X_{2}^{2}}{2}\right)=\sigma^{2}
$$

This is true even if $\left(X_{1}, X_{2}\right)$ are dependent. The estimate $\hat{\sigma}^{2}$ is unbiased.
3b. The distribution in question is the chi-square distribution with two degrees of freedom. From the table we find the $95 \%$ confidence interval to be

$$
I_{\sigma^{2}}=\left(0.27 \hat{\sigma}^{2}, 39.22 \hat{\sigma}^{2}\right)
$$

3c. It is easy to show that

$$
\tilde{\sigma}^{2}=s^{2}
$$

which implies that $\tilde{\sigma}^{2}$ is also an unbiased estimate of $\sigma^{2}$. It is not unbiased if ( $X_{1}, X_{2}$ ) are positively correlated, $\tilde{\sigma}^{2}$ would systematically underestimate $\sigma^{2}$.

3d. The variance $\tilde{\sigma}^{2}$ is twice as large as $\hat{\sigma}^{2}$, therefore $\hat{\sigma}^{2}$ is a better estimate.
4a. From the sample values we find

$$
\hat{x}_{0.25}=2, \quad \hat{x}_{0.5}=3.615, \quad \hat{x}_{0.75}=5.035
$$

which gives $\mathrm{IQR}=3.04$. Robustness means: if we add an unusually large sample value $s^{2}$ would change dramatically, but not IQR.

4b.

$$
\frac{3.04}{1.35}=2.25
$$

4c. The simplified normal probability plot is the scatter plot of three points on the plane $(2,-0.675),(3.615,0),(5.035,0.675)$.

5a. Pooling together 4000 observations for August 20019 we get

$$
\hat{p}_{2}=0.125, \quad s_{\hat{p}_{2}}=0.00523
$$

$5 b$. We apply the formula

$$
I_{p_{1}-p_{2}} \approx \hat{p}_{1}-\hat{p}_{2} \pm 1.96 \sqrt{\frac{\hat{p}_{1}\left(1-\hat{p}_{1}\right)}{n-1}+\frac{\hat{p}_{2}\left(1-\hat{p}_{2}\right)}{m-1}}
$$

which requires independence of two samples. Therefore we use only group 1 in 2018 and group 2 in 2019

$$
I_{p_{1}-p_{2}} \approx 0.10-0.13 \pm 1.96 \sqrt{\frac{0.1(1-0.1)}{1999}+\frac{0.13(1-0.13)}{1999}}=-0.03 \pm 0.02
$$

The interval does not cover zero and we conclude that the difference between $p_{1}$ and $p_{2}$ is significant at $5 \%$ level.

5 c . The matched pairs design results in the following observed counts

|  | 2019 yes | 2019 no | Total |
| :---: | :---: | :---: | :---: |
| 2018 yes | 120 | 80 | 200 |
| 2018 no | 120 | 1680 | 1800 |
| Total | 240 | 1760 | 2000 |

The McNemar's test statistic is

$$
\frac{(120-80)^{2}}{120+80}=8
$$

Taking the square root of 8 and using the normal distribution table we find the p -value

$$
2(1-0.9977)=0.005
$$

to be $0.5 \%$. We reject $H_{0}: p_{1}=p_{2}$.
6a. We apply a multiple regression setting

$$
Y=\beta_{0}+\beta_{1} x+\beta_{2} x^{2}+\epsilon
$$

where $\epsilon$ is a normally distributed $\mathrm{N}(0, \sigma)$ homoscedastic noise. Using the model utility test for

$$
H_{0}: \beta_{2}=0
$$

we observe that the corresponding confidence interval

$$
I_{\beta_{2}}=(-0.0017,-0.0015)
$$

does not cover 0 and we reject the null hypothesis at $5 \%$ significance level. We conclude that the quadratic model does appear to be suitable. To check the assumption concerning the noise one may draw a normal probability plot for the residuals.

6b. The adjusted coefficient of determination

$$
R_{a}^{2}=1-\frac{s^{2}}{s_{y}^{2}}=1-\frac{0.2874 / 5}{(5.3835)^{2}}=0.998
$$

says that $99.8 \%$ of the variation in the respons variable is explained by the quadratic model

$$
y=\beta_{0}+\beta_{1} x+\beta_{2} x^{2}
$$

6c. Because

$$
\sum_{i=1}^{n}\left(y_{i}-\hat{y}_{i}\right)=\sum_{i=1}^{n}\left(y_{i}-\bar{y}\right)+\sum_{i=1}^{n}\left(\hat{y}_{i}-\bar{y}\right)=s_{y} \sum_{i=1}^{n} \frac{\hat{y}_{i}-\bar{y}}{s_{y}}=s_{y} r \sum_{i=1}^{n} \frac{\hat{x}_{i}-\bar{x}}{s_{x}}=0
$$

6 d . The scatter plot of the data clearly indicates non-linear relationship - seemingly quadratic.

