## Tentamentsskrivning i Statistisk slutledning MVE155/MSG200, 7.5 hp .

Tid: 29 augusti 2019, kl 14.00-18.00
Examinator och jour: Serik Sagitov, tel. 031-772-5351, rum H3026 i MV-huset.
Hjälpmedel: Chalmersgodkänd räknare, egen formelsamling (fyra A4 sidor).
CTH: för " 3 " fordras 12 poäng, för " 4 " - 18 poäng, för " 5 " - 24 poäng.
GU : för " G " fordras 12 poäng, för "VG" - 20 poäng.
Inclusive eventuella bonuspoäng.

## Partial answers and solutions are also welcome. Good luck!

1. (5 points) A state teachers association studied the educational qualifications of its membership, which consists of 50000 teachers. The analyst used a sample design in which independent simple random samples of teachers are selected from among the member teachers in three school levels.

| Stratum | Level of school | Number of members | Sample size |
| :---: | :--- | :---: | :---: |
| 1 | Elementary | 25000 | 100 |
| 2 | High school | 20000 | 100 |
| 3 | College | 5000 | 100 |
|  | Total | 50000 | 300 |

(a) What are possible advantages and disadvantages of this sampling design compared to a single random sample of size $n=300$ taken from the whole population?
(b) The following data gives the sample means and variances for the reported numbers of years of education:

| Level of school | Mean | Variance |
| :--- | :---: | :---: |
| Elementary | 14.8 | 6.4 |
| High school | 17.3 | 2.7 |
| College | 19.3 | 3.6 |

Find a $99 \%$ confidence interval $I_{\mu}$ of the population mean. What are your assumptions?
(c) Without any prior knowledge on the variation within the strata, how would you allocate 300 observations among the three strata in a more effective way? Explain your choice.
2. (5 points) Turn to the data on the previous problem and treat the three strata as three separate populations. Denote by $\mu_{1}, \mu_{2}, \mu_{3}$ the corresponding population means.
(a) Build a $95 \%$ confidence interval for the difference $\delta=\mu_{1}-\mu_{3}$ based on a t-distribution. What is wrong with just checking if this interval does not contain zero, to reject the $H_{0}: \mu_{1}=\mu_{2}=\mu_{3}$ ?
(b) Test the null hypothesis of no difference $H_{0}: \mu_{1}=\mu_{2}=\mu_{3}$. What are your assumptions?
(c) A new study is planned to compare female and male means. There will be six independent samples of size 50 taken: so that for each of the level of school, 50 female members and 50 male members will be sampled at random. Describe a suitable parametric test.
3. (5 marks) Think of a chi-square distribution $\chi_{k}^{2}$ with $k$ degrees of freedom. It is connected to the standard normal distribution as follows: if $Z_{1}, \ldots, Z_{k}$ are $\mathrm{N}(0,1)$ and independent, then

$$
\begin{equation*}
Z_{1}^{2}+\ldots+Z_{k}^{2} \sim \chi_{k}^{2} \tag{1}
\end{equation*}
$$

Denote by $\mu$ and $\sigma^{2}$ the mean and variance of $\chi_{k}^{2}$-distribution.
(a) Using (1) compute $\mu$.
(b) Let $Z \sim \mathrm{~N}(0,1)$. Using the fact that the curtosis of a normal distribution equals 3 , show that

$$
\operatorname{Var}\left(Z^{2}\right)=2
$$

(c) Using (1) and (b) compute $\sigma^{2}$.
(d) The chi-square distribution has the same density as a gamma distribution with the shape parameter $\frac{k}{2}$ and the scale parameter $\frac{1}{2}$. Using this fact verify your answers for (a) and (c).
4. (5 points) Draw a copy of the boxplot on your answer paper.

(a) Compute the range and the inter-quartile range for the underlying sample.
(b) Is the distribution skewed to the right or to the left? Which is larger: the sample mean or the median? Explain.
(c) Roughly sketch the corresponding normal probability plot.
(d) Suggest a data transformation which will make the empirical distribution to look more like a normal distribution.
5. (5 points) Consider a cross-classification for a pair of categorical factors A and B. If factors $A$ and $B$ have three levels each, then the population distribution of a single cross classification outcome has the form

|  | $b_{1}$ | $b_{2}$ | $b_{3}$ | Total |
| :---: | :---: | :---: | :---: | :---: |
| $a_{1}$ | $\pi_{11}$ | $\pi_{12}$ | $\pi_{13}$ | $\pi_{1 .}$ |
| $a_{2}$ | $\pi_{21}$ | $\pi_{22}$ | $\pi_{23}$ | $\pi_{2 .}$ |
| $a_{3}$ | $\pi_{31}$ | $\pi_{I 2}$ | $\pi_{33}$ | $\pi_{3 .}$ |
| Total | $\pi_{\cdot 1}$ | $\pi_{\cdot 2}$ | $\pi_{\cdot 3}$ | 1 |

Here

$$
\pi_{i j}=\mathrm{P}\left(A=a_{i}, B=b_{j}\right)
$$

are the joint the probabilities, and

$$
\pi_{i}=\mathrm{P}\left(A=a_{i}\right), \quad \pi_{\cdot j}=\mathrm{P}\left(B=b_{j}\right)
$$

are the marginal probabilities.
(a) Clearly state the hypothesis of independence between factors A and B in terms of the above given parametric model.
(b) Present a realistic example of such two categorical factors. Try to make your example different from those mentioned in the lecture notes.
(c) Aiming at the chi-square test of independence, the following data was collected

|  | $b_{1}$ | $b_{2}$ | $b_{3}$ |
| :---: | :---: | :---: | :---: |
| $a_{1}$ | 6 | 3 | 9 |
| $a_{2}$ | 7 | 2 | 8 |
| $a_{3}$ | 13 | 6 | 15 |

Explain how these nine counts were produced. In particular, how many independent samples were collected?
(d) Apply the chi-square test of independence to the data in (c). What can be said about the p-value of the test?

6 (5 marks) An experiment was conducted to study the effects of two types of promotional expenditures on sales of a certain product sold in supermarkets. Sixteen localities were selected for the test. Different combinations of media advertising expenditures $\left(x_{1}\right)$ and point-of-sale expenditures $\left(x_{2}\right)$ were specified for the study, and the localities were assigned at random to each of these combinations $\left(x_{1 i}, x_{2 i}\right)$. Then the dollar sales $\left(y_{i}\right)$ were recorded for $i=1, \ldots, 16$.
(a) The sample correlation coefficient computed from the 16 pairs $\left(x_{1 i}, x_{2 i}\right)$ is zero. Is it a good or bad feature of the experimental design? Explain.
(b) The following table presents a part of the computer output for multiple regression

| Variable | Ref. coeff. | Std. dev. | T stat. |
| :--- | :---: | :---: | :---: |
| Constant | 2.13438 | .61036 | 3.50 |
| x1 | 3.02925 | .12028 | 25.18 |
| x2 | .70575 | .12028 | 5.87 |

Perform three utility tests. Clearly state your conclusions.
(c) The following table presents another part of the computer output for multiple regression

| Source | Sum of squares |
| :--- | ---: |
| Regression | 193.4888 |
| Residual | 3.7616 |
| Total | 197.2503 |

Compute the adjusted coefficient of determination. What does it say about the underlying model?
(d) For a supermarket with media advertising expenditures being $x_{1}=2$ and point-of-sale expenditures being $x_{2}=4$, what would be your prediction for the dollar sales of the product under the study?

Normal distribution table

| $z$ | 0.00 | 0.01 | 0.02 | 0.03 | 0.04 | 0.05 | 0.06 | 0.07 | 0.08 | 0.09 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.0 | 0.5000 | 0.5040 | 0.5080 | 0.5120 | 0.5160 | 0.5199 | 0.5239 | 0.5279 | 0.5319 | 0.5359 |
| 0.1 | 0.5398 | 0.5438 | 0.5478 | 0.5517 | 0.5557 | 0.5596 | 0.5636 | 0.5675 | 0.5714 | 0.5753 |
| 0.2 | 0.5793 | 0.5832 | 0.5871 | 0.5910 | 0.5948 | 0.5987 | 0.6026 | 0.6064 | 0.6103 | 0.6141 |
| 0.3 | 0.6179 | 0.6217 | 0.6255 | 0.6293 | 0.6331 | 0.6368 | 0.6406 | 0.6443 | 0.6480 | 0.6517 |
| 0.4 | 0.6554 | 0.6591 | 0.6628 | 0.6664 | 0.6700 | 0.6736 | 0.6772 | 0.6808 | 0.6844 | 0.6879 |
| 0.5 | 0.6915 | 0.6950 | 0.6985 | 0.7019 | 0.7054 | 0.7088 | 0.7123 | 0.7157 | 0.7190 | 0.7224 |
| 0.6 | 0.7257 | 0.7291 | 0.7324 | 0.7357 | 0.7389 | 0.7422 | 0.7454 | 0.7486 | 0.7517 | 0.7549 |
| 0.7 | 0.7580 | 0.7611 | 0.7642 | 0.7673 | 0.7704 | 0.7734 | 0.7764 | 0.7794 | 0.7823 | 0.7852 |
| 0.8 | 0.7881 | 0.7910 | 0.7939 | 0.7967 | 0.7995 | 0.8023 | 0.8051 | 0.8078 | 0.8106 | 0.8133 |
| 0.9 | 0.8159 | 0.8186 | 0.8212 | 0.8238 | 0.8264 | 0.8289 | 0.8315 | 0.8340 | 0.8365 | 0.8389 |
| 1.0 | 0.8413 | 0.8438 | 0.8461 | 0.8485 | 0.8508 | 0.8531 | 0.8554 | 0.8577 | 0.8599 | 0.8621 |
| 1.1 | 0.8643 | 0.8665 | 0.8686 | 0.8708 | 0.8729 | 0.8749 | 0.8770 | 0.8790 | 0.8810 | 0.8830 |
| 1.2 | 0.8849 | 0.8869 | 0.8888 | 0.8907 | 0.8925 | 0.8944 | 0.8962 | 0.8980 | 0.8997 | 0.9015 |
| 1.3 | 0.9032 | 0.9049 | 0.9066 | 0.9082 | 0.9099 | 0.9115 | 0.9131 | 0.9147 | 0.9162 | 0.9177 |
| 1.4 | 0.9192 | 0.9207 | 0.9222 | 0.9236 | 0.9251 | 0.9265 | 0.9279 | 0.9292 | 0.9306 | 0.9319 |
| 1.5 | 0.9332 | 0.9345 | 0.9357 | 0.9370 | 0.9382 | 0.9394 | 0.9406 | 0.9418 | 0.9429 | 0.9441 |
| 1.6 | 0.9452 | 0.9463 | 0.9474 | 0.9484 | 0.9495 | 0.9505 | 0.9515 | 0.9525 | 0.9535 | 0.9545 |
| 1.7 | 0.9554 | 0.9564 | 0.9573 | 0.9582 | 0.9591 | 0.9599 | 0.9608 | 0.9616 | 0.9625 | 0.9633 |
| 1.8 | 0.9641 | 0.9649 | 0.9656 | 0.9664 | 0.9671 | 0.9678 | 0.9686 | 0.9693 | 0.9699 | 0.9706 |
| 1.9 | 0.9713 | 0.9719 | 0.9726 | 0.9732 | 0.9738 | 0.9744 | 0.9750 | 0.9756 | 0.9761 | 0.9767 |
| 2.0 | 0.9772 | 0.9778 | 0.9783 | 0.9788 | 0.9793 | 0.9798 | 0.9803 | 0.9808 | 0.9812 | 0.9817 |
| 2.1 | 0.9821 | 0.9826 | 0.9830 | 0.9834 | 0.9838 | 0.9842 | 0.9846 | 0.9850 | 0.9854 | 0.9857 |
| 2.2 | 0.9861 | 0.9864 | 0.9868 | 0.9871 | 0.9875 | 0.9878 | 0.9881 | 0.9884 | 0.9887 | 0.9890 |
| 2.3 | 0.9893 | 0.9896 | 0.9898 | 0.9901 | 0.9904 | 0.9906 | 0.9909 | 0.9911 | 0.9913 | 0.9916 |
| 2.4 | 0.9918 | 0.9920 | 0.9922 | 0.9925 | 0.9927 | 0.9929 | 0.9931 | 0.9932 | 0.9934 | 0.9936 |
| 2.5 | 0.9938 | 0.9940 | 0.9941 | 0.9943 | 0.9945 | 0.9946 | 0.9948 | 0.9949 | 0.9951 | 0.9952 |
| 2.6 | 0.9953 | 0.9955 | 0.9956 | 0.9957 | 0.9959 | 0.9960 | 0.9961 | 0.9962 | 0.9963 | 0.9964 |
| 2.7 | 0.9965 | 0.9966 | 0.9967 | 0.9968 | 0.9969 | 0.9970 | 0.9971 | 0.9972 | 0.9973 | 0.9974 |
| 2.8 | 0.9974 | 0.9975 | 0.9976 | 0.9977 | 0.9977 | 0.9978 | 0.9979 | 0.9979 | 0.9980 | 0.9981 |
| 2.9 | 0.9981 | 0.9982 | 0.9982 | 0.9983 | 0.9984 | 0.9984 | 0.9985 | 0.9985 | 0.9986 | 0.9986 |
| 3.0 | 0.9987 | 0.9987 | 0.9987 | 0.9988 | 0.9988 | 0.9989 | 0.9989 | 0.9989 | 0.9990 | 0.9990 |
| 3.1 | 0.9990 | 0.9991 | 0.9991 | 0.9991 | 0.9992 | 0.9992 | 0.9992 | 0.9992 | 0.9993 | 0.9993 |
| 3.2 | 0.9993 | 0.9993 | 0.9994 | 0.9994 | 0.9994 | 0.9994 | 0.9994 | 0.9995 | 0.9995 | 0.9995 |
| 3.3 | 0.9995 | 0.9995 | 0.9995 | 0.9996 | 0.9996 | 0.9996 | 0.9996 | 0.9996 | 0.9996 | 0.9997 |
| 3.4 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9998 |

Chi-square distribution table


Area to the Right of the Critical Value of $\chi^{2}$

| df | 0.995 | 0.990 | 0.975 | 0.950 | 0.900 | 0.100 | 0.050 | 0.025 | 0.010 | 0.005 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 0.000 | 0.000 | 0.001 | 0.004 | 0.016 | 2.706 | 3.841 | 5.024 | 6.635 | 7.879 |
| 2 | 0.010 | 0.020 | 0.051 | 0.103 | 0.211 | 4.605 | 5.991 | 7.378 | 9.210 | 10.597 |
| 3 | 0.072 | 0.115 | 0.216 | 0.352 | 0.584 | 6.251 | 7.815 | 9.348 | 11.345 | 12.838 |
| 4 | 0.207 | 0.297 | 0.484 | 0.711 | 1.064 | 7.779 | 9.488 | 11.143 | 13.277 | 14.860 |
| 5 | 0.412 | 0.554 | 0.831 | 1.145 | 1.610 | 9.236 | 11.070 | 12.833 | 15.086 | 16.750 |
| 6 | 0.676 | 0.872 | 1.237 | 1.635 | 2.204 | 10.645 | 12.592 | 14.449 | 16.812 | 18.548 |
| 7 | 0.989 | 1.239 | 1.690 | 2.167 | 2.833 | 12.017 | 14.067 | 16.013 | 18.475 | 20.278 |
| 8 | 1.344 | 1.646 | 2.180 | 2.733 | 3.490 | 13.362 | 15.507 | 17.535 | 20.090 | 21.955 |
| 9 | 1.735 | 2.088 | 2.700 | 3.325 | 4.168 | 14.684 | 16.919 | 19.023 | 21.666 | 23.589 |
| 10 | 2.156 | 2.558 | 3.247 | 3.940 | 4.865 | 15.987 | 18.307 | 20.483 | 23.209 | 25.188 |
| 11 | 2.603 | 3.053 | 3.816 | 4.575 | 5.578 | 17.275 | 19.675 | 21.920 | 24.725 | 26.757 |
| 12 | 3.074 | 3.571 | 4.404 | 5.226 | 6.304 | 18.549 | 21.026 | 23.337 | 26.217 | 28.300 |
| 13 | 3.565 | 4.107 | 5.009 | 5.892 | 7.042 | 19.812 | 22.362 | 24.736 | 27.688 | 29.819 |
| 14 | 4.075 | 4.660 | 5.629 | 6.571 | 7.790 | 21.064 | 23.685 | 26.119 | 29.141 | 31.319 |
| 15 | 4.601 | 5.229 | 6.262 | 7.261 | 8.547 | 22.307 | 24.996 | 27.488 | 30.578 | 32.801 |
| 16 | 5.142 | 5.812 | 6.908 | 7.962 | 9.312 | 23.542 | 26.296 | 28.845 | 32.000 | 34.267 |
| 17 | 5.697 | 6.408 | 7.564 | 8.672 | 10.085 | 24.769 | 27.587 | 30.191 | 33.409 | 35.718 |
| 18 | 6.265 | 7.015 | 8.231 | 9.390 | 10.865 | 25.989 | 28.869 | 31.526 | 34.805 | 37.156 |
| 19 | 6.844 | 7.633 | 8.907 | 10.117 | 11.651 | 27.204 | 30.144 | 32.852 | 36.191 | 38.582 |
| 20 | 7.434 | 8.260 | 9.591 | 10.851 | 12.443 | 28.412 | 31.410 | 34.170 | 37.566 | 39.997 |
| 21 | 8.034 | 8.897 | 10.283 | 11.591 | 13.240 | 29.615 | 32.671 | 35.479 | 38.932 | 41.401 |
| 22 | 8.643 | 9.542 | 10.982 | 12.338 | 14.041 | 30.813 | 33.924 | 36.781 | 40.289 | 42.796 |
| 23 | 9.260 | 10.196 | 11.689 | 13.091 | 14.848 | 32.007 | 35.172 | 38.076 | 41.638 | 44.181 |
| 24 | 9.886 | 10.856 | 12.401 | 13.848 | 15.659 | 33.196 | 36.415 | 39.364 | 42.980 | 45.559 |
| 25 | 10.520 | 11.524 | 13.120 | 14.611 | 16.473 | 34.382 | 37.652 | 40.646 | 44.314 | 46.928 |
| 26 | 11.160 | 12.198 | 13.844 | 15.379 | 17.292 | 35.563 | 38.885 | 41.923 | 45.642 | 48.290 |
| 27 | 11.808 | 12.879 | 14.573 | 16.151 | 18.114 | 36.741 | 40.113 | 43.195 | 46.963 | 49.645 |
| 28 | 12.461 | 13.565 | 15.308 | 16.928 | 18.939 | 37.916 | 41.337 | 44.461 | 48.278 | 50.993 |
| 29 | 13.121 | 14.256 | 16.047 | 17.708 | 19.768 | 39.087 | 42.557 | 45.722 | 49.588 | 52.336 |
| 30 | 13.787 | 14.953 | 16.791 | 18.493 | 20.599 | 40.256 | 43.773 | 46.979 | 50.892 | 53.672 |
| 40 | 20.707 | 22.164 | 24.433 | 26.509 | 29.051 | 51.805 | 55.758 | 59.342 | 63.691 | 66.766 |
| 50 | 27.991 | 29.707 | 32.357 | 34.764 | 37.689 | 63.167 | 67.505 | 71.420 | 76.154 | 79.490 |
| 60 | 35.534 | 37.485 | 40.482 | 43.188 | 46.459 | 74.397 | 79.082 | 83.298 | 88.379 | 91.952 |
| 70 | 43.275 | 45.442 | 48.758 | 51.739 | 55.329 | 85.527 | 90.531 | 95.023 | 100.425 | 104.215 |
| 80 | 51.172 | 53.540 | 57.153 | 60.391 | 64.278 | 96.578 | 101.879 | 106.629 | 112.329 | 116.321 |
| 90 | 59.196 | 61.754 | 65.647 | 69.126 | 73.291 | 107.565 | 113.145 | 118.136 | 124.116 | 128.299 |
| 100 | 67.328 | 70.065 | 74.222 | 77.929 | 82.358 | 118.498 | 124.342 | 129.561 | 135.807 | 140.169 |

Critical values of $t$-distribution

| df/ $\boldsymbol{\alpha}=$ | .40 | .25 | .10 | .05 | .025 | .01 | .005 | .001 | .0005 |
| ---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | ---: |
| 1 | 0.325 | 1.000 | 3.078 | 6.314 | 12.706 | 31.821 | 63.657 | 318.309 | 636.619 |
| 2 | 0.289 | 0.816 | 1.886 | 2.920 | 4.303 | 6.965 | 9.925 | 22.327 | 31.599 |
| 3 | 0.277 | 0.765 | 1.638 | 2.353 | 3.182 | 4.541 | 5.841 | 10.215 | 12.924 |
| 4 | 0.271 | 0.741 | 1.533 | 2.132 | 2.776 | 3.747 | 4.604 | 7.173 | 8.610 |
| 5 | 0.267 | 0.727 | 1.476 | 2.015 | 2.571 | 3.365 | 4.032 | 5.893 | 6.869 |
| 6 | 0.265 | 0.718 | 1.440 | 1.943 | 2.447 | 3.143 | 3.707 | 5.208 | 5.959 |
| 7 | 0.263 | 0.711 | 1.415 | 1.895 | 2.365 | 2.998 | 3.499 | 4.785 | 5.408 |
| 8 | 0.262 | 0.706 | 1.397 | 1.860 | 2.306 | 2.896 | 3.355 | 4.501 | 5.041 |
| 9 | 0.261 | 0.703 | 1.383 | 1.833 | 2.262 | 2.821 | 3.250 | 4.297 | 4.781 |
| 10 | 0.260 | 0.700 | 1.372 | 1.812 | 2.228 | 2.764 | 3.169 | 4.144 | 4.587 |
| 11 | 0.260 | 0.697 | 1.363 | 1.796 | 2.201 | 2.718 | 3.106 | 4.025 | 4.437 |
| 12 | 0.259 | 0.695 | 1.356 | 1.782 | 2.179 | 2.681 | 3.055 | 3.930 | 4.318 |
| 13 | 0.259 | 0.694 | 1.350 | 1.771 | 2.160 | 2.650 | 3.012 | 3.852 | 4.221 |
| 14 | 0.258 | 0.692 | 1.345 | 1.761 | 2.145 | 2.624 | 2.977 | 3.787 | 4.140 |
| 15 | 0.258 | 0.691 | 1.341 | 1.753 | 2.131 | 2.602 | 2.947 | 3.733 | 4.073 |
| 16 | 0.258 | 0.690 | 1.337 | 1.746 | 2.120 | 2.583 | 2.921 | 3.686 | 4.015 |
| 17 | 0.257 | 0.689 | 1.333 | 1.740 | 2.110 | 2.567 | 2.898 | 3.646 | 3.965 |
| 18 | 0.257 | 0.688 | 1.330 | 1.734 | 2.101 | 2.552 | 2.878 | 3.610 | 3.922 |
| 19 | 0.257 | 0.688 | 1.328 | 1.729 | 2.093 | 2.539 | 2.861 | 3.579 | 3.883 |
| 20 | 0.257 | 0.687 | 1.325 | 1.725 | 2.086 | 2.528 | 2.845 | 3.552 | 3.850 |
| 21 | 0.257 | 0.686 | 1.323 | 1.721 | 2.080 | 2.518 | 2.831 | 3.527 | 3.819 |
| 22 | 0.256 | 0.686 | 1.321 | 1.717 | 2.074 | 2.508 | 2.819 | 3.505 | 3.792 |
| 23 | 0.256 | 0.685 | 1.319 | 1.714 | 2.069 | 2.500 | 2.807 | 3.485 | 3.768 |
| 24 | 0.256 | 0.685 | 1.318 | 1.711 | 2.064 | 2.492 | 2.797 | 3.467 | 3.745 |
| 25 | 0.256 | 0.684 | 1.316 | 1.708 | 2.060 | 2.485 | 2.787 | 3.450 | 3.725 |
| 26 | 0.256 | 0.684 | 1.315 | 1.706 | 2.056 | 2.479 | 2.779 | 3.435 | 3.707 |
| 27 | 0.256 | 0.684 | 1.314 | 1.703 | 2.052 | 2.473 | 2.771 | 3.421 | 3.690 |
| 28 | 0.256 | 0.683 | 1.313 | 1.701 | 2.048 | 2.467 | 2.763 | 3.408 | 3.674 |
| 29 | 0.256 | 0.683 | 1.311 | 1.699 | 2.045 | 2.462 | 2.756 | 3.396 | 3.659 |
| 30 | 0.256 | 0.683 | 1.310 | 1.697 | 2.042 | 2.457 | 2.750 | 3.385 | 3.646 |
| 35 | 0.255 | 0.682 | 1.306 | 1.690 | 2.030 | 2.438 | 2.724 | 3.340 | 3.591 |
| 40 | 0.255 | 0.681 | 1.303 | 1.684 | 2.021 | 2.423 | 2.704 | 3.307 | 3.551 |
| 50 | 0.255 | 0.679 | 1.299 | 1.676 | 2.009 | 2.403 | 2.678 | 3.261 | 3.496 |
| 60 | 0.254 | 0.679 | 1.296 | 1.671 | 2.000 | 2.390 | 2.660 | 3.232 | 3.460 |
| 120 | 0.254 | 0.677 | 1.289 | 1.658 | 1.980 | 2.358 | 2.617 | 3.160 | 3.373 |
| inf. | 0.253 | 0.674 | 1.282 | 1.645 | 1.960 | 2.326 | 2.576 | 3.090 | 3.291 |
|  |  |  |  |  |  |  |  |  |  |

Critical values of the F-distribution (continued)

|  |  |  |  |  |  | Degrees of freedom in the numerator |  |  |  |  |
| :---: | :---: | ---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | 1 | 2 | 3 | 4 | 5 | 6 |

## NUMERICAL ANSWERS

1a. An advantage of this stratified sampling compared to a single sample is in the possibility of comparing qualification levels of the teachers from different school categories. A possible disadvantage is a larger standard error of the estimate of the overall mean.

1 b . For the stratified sample mean with $w_{1}=0.5, w_{2}=0.4, w_{3}=0.1$, and $n_{1}=n_{2}=n_{3}=100$, we find the stratified sample mean

$$
\bar{x}_{\mathrm{s}}=0.5 \cdot 14.8+0.4 \cdot 17.3+0.1 \cdot 19.3=16.25
$$

and its standard error

$$
s_{\bar{x}_{\mathrm{s}}}=\sqrt{\frac{(0.5)^{2} 6.4}{100}+\frac{(0.4)^{2} 2.7}{100}+\frac{(0.1)^{2} 3.6}{100}}=0.144
$$

The $99 \%$ confidence interval for the overall mean becomes

$$
I_{\mu}=16.25 \pm 2.58 \cdot 0.144=16.25 \pm 0.37=(15.88,16.62)
$$

1c. Using proportional allocation allocation, we put 150 observations on the Elementary level, 120 observations on the High school level, and 30 observations on the College level. This would make smaller standard error compared to the simple random sample.

2a. Two sample $95 \%$ confidence interval

$$
I_{\delta}=14.8-19.3 \pm 1.96 \sqrt{\frac{3.6}{100}+\frac{6.4}{100}}=-4.50 \pm 0.62
$$

Since this interval does not cover zero, it indicates that the null hypothesis $H_{0}: \mu_{1}=\mu_{2}=\mu_{3}$ should be rejected. However, this conclusion is based on the comparison of the two extreme strata means and might be false due to missing of the multiple comparison effect.

2b. We apply the one-way anova test. Using the the samples means and variances we obtain

$$
\begin{aligned}
& \frac{14.8+17.3+19.3}{3}=17.13 \\
& S S_{A}=100 \cdot\left((14.8-17.13)^{2}+(17.3-17.13)^{2}+(19.3-17.13)^{2}\right)=1016.7 \\
& S S_{E}=6.4 \cdot 99+2.7 \cdot 99+3.6 \cdot 99=1257.3
\end{aligned}
$$

Putting these into the anova table

| Source of variation | SS | df | MS | F |
| :--- | :---: | :---: | :---: | :---: |
| Main factor | 1016.7 | 2 | 508.3 | 121.0 |
| Error | 1257.3 | 297 | 4.2 |  |
| Total | 2274.0 | 299 |  |  |

and checking the F-distribution table with 2 degrees of freedom in the numerator, we see that the p-value of the F-test is much less than $0.1 \%$. We conclude that the difference between the three means is statistically significant.

2c. A suitable parametric model would be a two-way anova model with two main factors being factor $\mathrm{A}=$ level of school with 3 levels, factor $B=$ gender with 2 levels.

3a. Since

$$
\mathrm{E}\left(Z^{2}\right)=\operatorname{Var}(Z)=1
$$

we get

$$
\mu=\mathrm{E}\left(Z_{1}^{2}+\ldots+Z_{k}^{2}\right)=k
$$

3b. Kurtosis of the standard normal distribution is

$$
3=\mathrm{E}\left((Z)^{4}\right)
$$

and therefore,

$$
\operatorname{Var}\left(Z^{2}\right)=\mathrm{E}\left((Z)^{4}\right)-\left(\mathrm{E}\left(Z^{2}\right)\right)^{2}=2
$$

3c. Due to independence,

$$
\operatorname{Var}\left(Z_{1}^{2}+\ldots+Z_{k}^{2}\right)=k \operatorname{Var}\left(Z^{2}\right)=2 k
$$

3d. Gamma distribution $\operatorname{Gam}(\alpha, \lambda)$ has the mean and variance

$$
\mu=\frac{\alpha}{\lambda}, \quad \sigma^{2}=\frac{\alpha}{\lambda^{2}} .
$$

With $\alpha=\frac{k}{2}$ and $\mu=\frac{1}{2}$, we obtain the same values as above

$$
\mu=\frac{k / 2}{1 / 2}=k, \quad \sigma^{2}=\frac{k / 2}{1 / 4}=2 k .
$$

4a. The range is close to 3500 and the inter-quartile range is like 500 .


4b. Skewed the right. The sample mean is larger than the sample median as the sample values larger than the median contribute to the arithmetic mean such that the mean will be larger than the median.

4c. For a right skewed distribution the normal probability plot would have the following typical profile reflecting the fact that the largest sample values are larger than would be predicted by the normal distribution.


4d. A logarithmic transformation

$$
\text { new data }=\log (\text { original data })
$$

might do the job.
5a. Hypothesis of independence

$$
\pi_{i j}=\pi_{i} \cdot \pi_{\cdot j}, \quad i=1,2,3, \quad j=1,2,3
$$

5c. The nine numbers

|  | $b_{1}$ | $b_{2}$ | $b_{3}$ |
| :---: | :---: | :---: | :---: |
| $a_{1}$ | 6 | 3 | 9 |
| $a_{2}$ | 7 | 2 | 8 |
| $a_{3}$ | 13 | 6 | 15 |

are the counts obtained after cross-classification of a single sample.
5 d . The observed and expected counts

|  | $b_{1}$ | $b_{2}$ | $b_{3}$ |
| :---: | :---: | :---: | :---: |
| $a_{1}$ | $6(6.8)$ | $3(2.9)$ | $9(8.3)$ |
| $a_{2}$ | $7(6.4)$ | $2(2.7)$ | $8(7.9)$ |
| $a_{3}$ | $13(12.8)$ | $6(5.4)$ | $15(15.8)$ |

Here the problems is with the expected counts being smaller than recommended 5 . Combining the first two rows we get a new table with observed and expected counts

|  | $b_{1}$ | $b_{2}$ | $b_{3}$ |
| :---: | :---: | :---: | :---: |
| $a_{1}+a_{2}$ | $13(13.2)$ | $5(5.6)$ | $17(16.2)$ |
| $a_{3}$ | $13(12.8)$ | $6(5.4)$ | $15(15.8)$ |

The corresponding chi-square test statistic is as small as 0.2 , and according to the $\chi_{2}^{2}$-distribution table the p-value is close to $90 \%$. Thus we do not reject the null hypothesis of independence.

6a. A problematic case in a multiple regression setting

$$
Y=\beta_{0}+\beta_{1} X_{1}+\beta_{2} X_{2}+\epsilon
$$

is when two explanatory variables are linearly dependent (collinearity). In our case the $X_{1}$ and $X_{2}$ are uncorrelated (orthogonality) which is a very good feature for the multiple regression design of experiment.

6b. Using $t_{13}$ as the null distribution for all three test statistics for three utility tests we reject the following three null hypotheses

$$
H_{0}: \beta_{0}=0, \quad H_{0}: \beta_{1}=0, \quad H_{0}: \beta_{2}=0
$$

$6 c$. The adjusted coefficient of determination

$$
R_{a}^{2}=1-\frac{15 \cdot 3.76}{13 \cdot 197.25}=0.98
$$

gives a high score ( 98 out of 100) on how well the two explanatory variables explain the observed variation in the response variable.

6 d . The predicted mean response is

$$
2.13+3.03 \cdot 2+0.71 \cdot 4=11.3
$$

