Examination for the course Foundations of Probability Theory.

Friday, 21 December 2007, 08.30-13.30 in the H house.

Examiner: Torgny Lindvall. Telephone connect. 3574 or mobile 0705-987486.

Teacher available at the examination site around 10.00 and 11.45.

Facilities: Dictionaries, from and into English.

A completely solved problem gives 5 credit points.

We suppose that events and random variables are defined on a probability space that we call  $(\Omega, \mathcal{F}, \mathbf{P})$ .

- 1. The random variables X and Y are independent and have Poisson distributions. Use probability generating functions to find the distribution of X + Y.
- 2. We have an infinite sequence of events  $A_1, A_2, \ldots$  Prove Boole's Inequality:  $\mathbf{P}(\bigcup_{1}^{\infty} A_k) \leq \sum_{1}^{\infty} \mathbf{P}(A_k)$ .
- 3. Consider a symmetric Simple Random Walk  $\mathbf{W} = (W_n)_0^{\infty}$  starting at an integer a > 0:  $W_n = a + \sum_{1}^{n} X_k$ , where the  $X_k$ 's are the iid steps;  $\mathbf{P}(X_k = 1) = \mathbf{P}(X_k = -1) = \frac{1}{2}$ . What is the probability  $\mathbf{P}(\mathbf{W})$  visits b before 0) where b is an integer > a? Hint: establish a difference equation; it is easily solved in the symmetric case!
- 4. The random variables N and  $X_1, X_2, \ldots$  are independent and have finite expectations and variances, the  $X_k$ -variables are identically distributed and have expectation = 0, and N takes values in  $\mathcal{Z}_+$ . Let  $S_n = \sum_1^n X_k$  for  $n \in \mathcal{Z}_+$  ( $S_0 = 0$ ). Make a careful calculation of  $Var[S_N]$  (every step commented!) in terms of the expectations and variances of N and  $X_k$ .
- 5. Give an example of two non-negative random variables X and Y with the same distribution such that  $\mathbf{E}[X] < \infty$  but  $\mathbf{E}[X \cdot Y] = \infty$ .
- 6. Let  $X_1, X_2, \ldots$  be iid with  $E[X_1] = 0$  and  $E[X_1^4] < \infty$ . Formulate and prove the Strong Law of Large Numbers for  $X_1, X_2, \ldots$

FOR SHORT SOLUTIONS: PLEASE TURN PAGE!

Short solutions to Foundations of Probability Theory 21 Dec 2007. Examiner: Torgny Lindvall.

- 1. Cf. Williams [W], 5.2, pp. 143-144. Suppose X and Y have parameters  $\lambda$  and  $\mu$  respectively. Due to the independence,  $g_{X+Y}(s) = g_X(s) \cdot g_Y(s) = \exp(\lambda(s-1)) \cdot \exp(\mu(s-1)) = \exp((\lambda+\mu)(s-1))$  for  $0 \le s \le 1$ , which is the pgf of the Poisson distribution with parameter  $\lambda + \mu$ . Since a pgf determines the distribution on  $\mathcal{Z}_+$ , we have solved the problem.
- 2. Let  $B_1 = A_1$ , and for  $k \geq 2$ , let  $B_k = A_k \setminus \bigcup_{1}^{k-1} A_i$ . We have  $\bigcup_{1}^{\infty} A_k = \bigcup_{1}^{\infty} B_k$ . Since  $B_k \subseteq A_k$  for all  $k \geq 1$  and the  $B_k$ 's are disjoint, we obtain  $\mathbf{P}(\bigcup_{1}^{\infty} A_k) = \mathbf{P}(\bigcup_{1}^{\infty} B_k) = \sum_{1}^{\infty} \mathbf{P}(B_k) \leq \sum \mathbf{P}(A_k)$ .
- 3. Cf. [W], 4.4, pp. 117-118. With the notation used there, we have  $x(k) = \frac{1}{2}(x(k-1)+x(k+1))$  for  $1 \le k \le b-1$ , which implies that x(k+1)-x(k) = x(k)-x(k-1) for those values of k. It is now easy to see that x(k) = k/b for all  $0 \le k \le b$  since x(0) = 0 and x(b) = 1; hence the asked for probability equals a/b.
- 4. Cf. [W], 9.1, pp. 392-393. Use the arguments there to prove that  $\mathbf{E}[S_N] = 0$ . This means that  $\mathbf{Var}[S_N] = \mathbf{E}[S_N^2]$ . Using analogous arguments, we get  $\mathbf{E}[S_N^2] = \sum_0^\infty \mathbf{E}[S_N^2|N=k] \cdot \mathbf{P}(N=k) = \sum_0^\infty \mathbf{E}[S_k^2] \cdot \mathbf{P}(N=k) = \sum_0^\infty \mathbf{Var}[S_k] \cdot \mathbf{P}(N=k) = \mathbf{Var}[X_1] \cdot \sum_0^\infty k \cdot \mathbf{P}(N=k) = \mathbf{Var}[X_1] \cdot \mathbf{E}[N]$ .
- 5. We let  $X = \exp((3/4) \cdot X')$ , where X' is exponentially distributed with rate  $\lambda = 1$ . Then  $\mathbf{E}[X] = \int_0^\infty \exp((3/4) \cdot x) \cdot \exp(-x) dx < \infty$ , but if we let Y = X, then  $\mathbf{E}[XY] = \mathbf{E}[X^2] = \infty$ . Alternatively, let X have density  $f(x) = C/(1+x^3)$  for  $x \geq 0$  where C is a normalizing constant. Again:  $\mathbf{E}[X] < \infty$  but  $\mathbf{E}[X^2] = \infty$ .
- 6. Cf. [W], 4.3, pp. 113-114.